



A Symmetric Information Divergence Measure of the Csiszár's f -Divergence Class and Its Bounds

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Abstract—A symmetric measure of information divergence is proposed. This measure belongs to the class of Csiszár's f -divergences. Its properties are studied and bounds in terms of well known divergence measures obtained. A numerical illustration is carried out to compare this measure with some known divergence measures. © 2005 Elsevier Ltd. All rights reserved.

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1. INTRODUCTION

There are several types of information divergence measures studied in literature which compare two probability distributions and have applications in information theory, statistics and engineering. A convenient classification to differentiate these measures is to categorize them as parametric, nonparametric and entropy-type measures of information [1]. Parametric measures of information measure the amount of information about an unknown parameter θ supplied by the data and are functions of θ . The best known measure of this type is Fisher's measure of information [2]. Nonparametric measures give the amount of information supplied by the data for discriminating in favor of a probability distribution f_1 against another f_2 , or for measuring the distance or affinity between f_1 and f_2 . The Kullback-Leibler measure is the best known in this class [3]. Measures of entropy express the amount of information contained in a distribution, that is, the amount of uncertainty associated with the outcome of an experiment. The classical measures of this type are Shannon's [4] and Rényi's measures [5]. The construction of measures of information divergence is not an easy task. Methods for deriving parametric measures of information from the nonparametric ones and from the information matrices are suggested in [1].

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In this paper, we present a new symmetric nonparametric information divergence measure which belongs to the class of Csiszár's f -divergences [6,7]. In Section 2, we discuss the Csiszár's f -divergences and their properties. Information inequalities are presented in Section 3. New symmetric divergence measure and its bounds are obtained in Section 4. In Section 5, it is shown that the suggested measure can be applied to the parametric family of distributions. A numerical illustration for studying the behavior of new measure is shown in Section 6. Section 7 concludes the paper.

2. CSISZÁR'S F -DIVERGENCES AND PROPERTIES

Let $\Omega = \{x_1, x_2, \dots\}$ be a set with at least two elements, $\beta(\Omega)$, the set of all subsets of Ω and \mathbb{P} , the set of all probability distributions $P = (p(x) : x \in \Omega)$ on Ω . A pair $(P, Q) \in \mathbb{P}^2$ of probability distributions is called a *simple versus simple testing problem*. Two probability distributions P and Q are called *orthogonal* ($P \perp Q$) if there exists an element $A \in \beta(\Omega)$ such that $P(A) = Q(A^c) = 0$, where $A^c = \Omega/A$. A testing problem $(P, Q) \in \mathbb{P}^2$ is called *least informative* if $P = Q$ and *most informative* if $P \perp Q$. Further, let \mathbb{F} be a set of convex functions $f : [0, \infty) \rightarrow (-\infty, \infty]$ continuous at 0, that is, $f(0) = \lim_{u \downarrow 0} f(u)$, $\mathbb{F}_0 = \{f \in \mathbb{F} : f(1) = 0\}$ and let D_-f and D_+f denote the *left-hand side* and *right-hand side derivatives* of f , respectively. Define $f^* \in \mathbb{F}$, the **-conjugate* (convex) function of f , by $f^*(u) = uf(1/u)$, $u \in (0, \infty)$ and $\tilde{f} = f + f^*$. For a convex function $f : [0, \infty) \rightarrow \mathbb{R}$, the *f -divergence* of the probability distributions P and Q is defined [6–8],

$$C_f(P, Q) = \sum_{x \in \Omega} q(x) f\left(\frac{p(x)}{q(x)}\right). \quad (2.1)$$

It is well known that $C_f(P, Q)$ is a versatile functional form which results in a number of popular divergence measures [9,10]. Most common choices of f satisfy $f(1) = 0$, so that $C_f(P, P) = 0$. Convexity ensures that divergence measure $C_f(P, Q)$ is nonnegative. Some examples are $f(u) = u \ln u$ provides the Kullback-Leibler measure [3], $f(u) = |u - 1|$ results in the variational distance [11,12], $f(u) = (u - 1)^2$ yields the χ^2 -divergence [13].

The basic general properties of f -divergences including their axiomatic properties and some important classes are given in [9]. For $f, f^*, f_1 \in \mathbb{F}, \forall (P, Q) \in \mathbb{P}^2, u \in (0, \infty)$,

- (i) $C_f(P, Q) = C_{f^*}(Q, P)$.
- (ii) UNIQUENESS THEOREM. (See [14].)

$$I_{f_1}(P, Q) = I_f(P, Q), \quad \text{iff } \exists c \in \mathbb{R} : f_1(u) - f(u) = c(u - 1).$$

- (iii) Let $c \in [D_-f(1), D_+f(1)]$. Then, $f_1(u) = f(u) - c(u - 1)$ satisfies $f_1(u) \geq f(1), \forall u \in [0, \infty)$ while not changing the f -divergence. Hence, without loss of generality $f_1(u) \geq f(1), \forall u \in [0, \infty)$.
- (iv) SYMMETRY THEOREM. (See [14].)

$$I_{f^*}(P, Q) = I_f(P, Q), \quad \text{iff } \exists c \in \mathbb{R} : f^*(u) - f(u) = c(u - 1).$$

- (v) RANGE OF VALUES THEOREM. (See [15].)

$$f(1) \leq I_f(P, Q) \leq f(0) + f^*(0).$$

In the first inequality, equality holds iff $P = Q$. The latter provides f is strictly convex at 1. The difference $I_f(P, Q) - f(1)$ is a quantity that compares the given testing problem $(P, Q) \in \mathbb{P}^2$ with the *least* informative testing problem. Given $f \in \mathbb{F}$, by setting $f(u) := f(u) - f(1)$, we can have $f(1) = 0$ and hence, without loss of generality, $f \in \mathbb{F}_0$. Thus, $I_f(P, Q)$ serves as an appropriate *measure of similarity* between two distributions.

In the second inequality, equality holds iff $P \perp Q$. The latter provides $\bar{f}(0) \leq f(0) + f^*(0) < \infty$. The difference $I_g(P, Q) := \bar{f}(0) - I_f(P, Q)$ is a quantity that compares the given testing problem $(P, Q) \in \mathbb{P}^2$ with the *most* informative testing problem. Thus, $I_g(P, Q)$ serves as an appropriate *measure of orthogonality* for the two distributions where the concave function $g : [0, \infty) \rightarrow \mathbb{R}$ is given by $g(u) = f(0) + uf^*(0) - f(u)$.

- (vi) CHARACTERIZATION THEOREM. (See [7].) Given a mapping $I : \mathbb{P}^2 \mapsto (-\infty, \infty)$,
 - (a) I is an f -divergence, that is, there exists an $f \in \mathbb{F}$, such that

$$I(P, Q) = C_f(P, Q), \quad \forall (P, Q) \in \mathbb{P}^2.$$

- (b) $C_f(P, Q)$ is invariant under permutation of Ω .
- (c) Let $\mathbb{A} = (A_i, i \geq 1)$ be a partition of Ω , and $P_{\mathbb{A}} = (P(A_i), i \geq 1)$ and $Q_{\mathbb{A}} = (Q(A_i), i \geq 1)$ be the restrictions of the probability distributions P and Q to \mathbb{A} . Then, $I(P, Q) \geq I(P_{\mathbb{A}}, Q_{\mathbb{A}})$ with equality if $P(A_i) \times p(x) = Q(A_i) \times q(x), \forall x \in A_i, i \geq 1$.
- (d) Let P_1, P_2 and Q_1, Q_2 be probability distributions on Ω . Then,

$$I(\alpha P_1 + (1 - \alpha) P_2, \alpha Q_1 + (1 - \alpha) Q_2) \leq \alpha I(P_1, Q_1) + (1 - \alpha) I(P_2, Q_2).$$

By characterization theorem, the *-conjugate of a convex function f is $f^*(u) \equiv uf(1/u)$.

For brevity, in what follows now, we will denote $C_f(P, Q), p(x), q(x)$, and $\sum_{x \in \Omega}$ by $C(P, Q), p, q$, and \sum , respectively.

Some popularly practised information divergence measures are as follows.

χ^α -Divergences.

VARIATIONAL DISTANCE. (See [11,12].)

$$V(P, Q) = \sum |p - q|. \tag{2.2}$$

χ^2 -DIVERGENCE. (See [13].)

$$\chi^2(P, Q) = \sum \frac{(p - q)^2}{q} = \sum \frac{p^2}{q} - 1. \tag{2.3}$$

SYMMETRIC χ^2 -DIVERGENCE.

$$\Psi(P, Q) = \chi^2(P, Q) + \chi^2(Q, P) = \sum \frac{(p + q)(p - q)^2}{pq}. \tag{2.4}$$

KULLBACK AND LEIBLER. (See [3].)

$$K(P, Q) = \sum p \ln \left(\frac{p}{q} \right). \tag{2.5}$$

KULLBACK-LEIBLER SYMMETRIC DIVERGENCE.

$$J(P, Q) = K(P, Q) + K(Q, P) = \sum (p - q) \ln \left(\frac{p}{q} \right). \tag{2.6}$$

TRIANGULAR DISCRIMINATION. (See [16,17].)

$$\Delta(P, Q) = \sum \frac{|p - q|^2}{p + q}. \tag{2.7}$$

SIBSON INFORMATION RADIUS. (See [18–21].)

$$I_r(P, Q) = \begin{cases} (r-1)^{-1} \left[\sum \left(\frac{p^r + q^r}{2} \right) \left(\frac{p+q}{2} \right)^{1-r} - 1 \right], & r \neq 1, r > 0, \\ \sum \frac{p \ln p + q \ln q}{2} - \left(\frac{p+q}{2} \right) \ln \left(\frac{p+q}{2} \right), & r = 1. \end{cases} \quad (2.8)$$

TANEJA DIVERGENCE MEASURE. (See [22].)

$$T_r(P, Q) = \begin{cases} (r-1)^{-1} \left[\sum \left(\frac{p^{1-r} + q^{1-r}}{2} \right) \left(\frac{p+q}{2} \right)^r - 1 \right], & r \neq 1, r > 0, \\ \sum \left(\frac{p+q}{2} \right) \ln \left(\frac{p+q}{2\sqrt{pq}} \right), & r = 1. \end{cases} \quad (2.9)$$

The following divergences are famous divergence measures. It may be noted that they are not members of the family of Csiszár's f -divergences (since the functions $g^\alpha(u) = u^\alpha$, $\alpha \in (0, 1)$ are concave).

BHATTACHARYYA DISTANCE. (See [23].)

$$B(P, Q) = \sum \sqrt{pq}. \quad (2.10)$$

HELLINGER DISCRIMINATION. (See [24].)

$$h(P, Q) = \sum \frac{(\sqrt{p} - \sqrt{q})^2}{2}. \quad (2.11)$$

RÉNYI MEASURE. (See [5].)

$$R_r(P, Q) = \begin{cases} (r-1)^{-1} \ln \left(\sum p^r q^{1-r} \right), & r \in (0, \infty) \setminus \{1\}, \\ \sum p \ln \frac{p}{q}, & r = 1. \end{cases} \quad (2.12)$$

We have the following inequalities relating to $V(P, Q)$, $\Delta(P, Q)$, $K(P, Q)$, and $h(P, Q)$.

CSISZÁR. (See [6].)

$$K(P, Q) \geq \frac{V^2(P, Q)}{2}. \quad (2.13)$$

CSISZÁR. (See [6,7].)

$$K(P, Q) \geq \frac{V^2(P, Q)}{2} + \frac{V^4(P, Q)}{36}. \quad (2.14)$$

TOPSØE. (See [17].)

$$K(P, Q) \geq \frac{V^2(P, Q)}{2} + \frac{V^4(P, Q)}{36} + \frac{V^6(P, Q)}{270} + \frac{V^8(P, Q)}{340200}. \quad (2.15)$$

VAJDA AND TOUSSAINT. (See [15] and [25], respectively.)

$$K(P, Q) \geq \max \{L_1(V), L_2(V)\}, \quad (2.16)$$

where, from [15],

$$L_1(V) = \ln \left(\frac{2+V}{2-V} \right) - \frac{2V}{2+V}, \quad 0 \leq V \leq 2,$$

and from [26],

$$L_2(V) = \frac{V^2}{2} + \frac{V^4}{36} + \frac{V^8}{288}, \quad 0 \leq V \leq 2.$$

TOPSØE. (See [18].)

$$\frac{1}{2}V^2(P, Q) \leq \Delta(P, Q) \leq V(P, Q). \tag{2.17}$$

LECAM AND DACUNHA-CASTELLE (SEE [16] AND [17], RESPECTIVELY.).

$$2h(P, Q) \leq \Delta(P, Q) \leq 4h(P, Q). \tag{2.18}$$

KRAFT. (See [27].)

$$\frac{1}{8}V^2(P, Q) \leq h(P, Q) \left(1 - \frac{1}{2}h(P, Q)\right). \tag{2.19}$$

TOPSØE. (See [18].)

$$\frac{1}{8}V^2(P, Q) \leq h(P, Q) \leq \frac{1}{2}V(P, Q), \tag{2.20}$$

and

$$K(P, Q) \leq (\log 2)V(P, Q) + \log c, \tag{2.21}$$

where $c = \max(p_i/q_i), \forall i = 1, \dots, n$.

3. INFORMATION BOUNDS

Different kinds of bounds on the distance, information and divergence measures have been studied recently [28–34]. In [28], authors unified and generalized information bounds for $C(P, Q)$ studied in [29–34] given in the following theorem.

THEOREM 3. *Let $f : I \subset \mathbb{R}_+ \rightarrow \mathbb{R}$ be a mapping which is normalized, i.e., $f(1) = 0$ and suppose that*

- (i) *f is twice differentiable on $(r, R), 0 \leq r \leq 1 \leq R < \infty$, (f' and f'' denote the first and second derivatives of f),*
- (ii) *there exists real constants m, M , such that $m < M$ and $m \leq x^{2-s}f''(x) \leq M, \forall x \in (r, R), s \in \mathbb{R}$.*

If $P, Q \in \mathbb{P}^2$ are discrete probability distributions with $0 < r \leq p/q \leq R < \infty$,

$$m\Phi_s(P, Q) \leq C(P, Q) \leq M\Phi_s(P, Q), \tag{3.1}$$

and

$$m(\eta_s(P, Q) - \Phi_s(P, Q)) \leq C_\rho(P, Q) - C(P, Q) \leq M(\eta_s(P, Q) - \Phi_s(P, Q)), \tag{3.2}$$

where

$$\Phi_s(P, Q) = \begin{cases} {}^2K_s(P, Q), & s \neq 0, 1, \\ K(Q, P), & s = 0, \\ K(P, Q), & s = 1, \end{cases} \tag{3.3}$$

$${}^2K_s(P, Q) = [s(s-1)]^{-1} \left[\sum p^s q^{1-s} - 1 \right], \quad s \neq 0, 1, \tag{3.4}$$

$$K(P, Q) = \sum p \ln \left(\frac{p}{q} \right),$$

$$C_\rho(P, Q) = C_{f'} \left(\frac{P^2}{Q}, P \right) - C_{f'}(P, Q) = \sum (p - q) f' \left(\frac{p}{q} \right), \tag{3.5}$$

$$\eta_s(P, Q) = C_{\phi'_s} \left(\frac{P^2}{Q}, P \right) - C_{\phi'_s}(P, Q),$$

$$= \begin{cases} (s-1)^{-1} \sum (p - q) \left(\frac{p}{q} \right)^{s-1}, & s \neq 1, \\ \sum (p - q) \ln \left(\frac{p}{q} \right), & s = 1 \end{cases} \tag{3.6}$$

As a consequence of this theorem, following information inequalities which are interesting from the *information-theoretic* point of view, are also obtained in [28].

(i) The case $s = 2$ provides the information bounds in terms of the χ^2 -divergence, $\chi^2(P, Q)$,

$$\frac{m}{2}\chi^2(P, Q) \leq C(P, Q) \leq \frac{M}{2}\chi^2(P, Q), \quad (3.7)$$

and

$$\frac{m}{2}\chi^2(P, Q) \leq C_\rho(P, Q) - C(P, Q) \leq \frac{M}{2}\chi^2(P, Q). \quad (3.8)$$

(ii) For $s = 1$, the information bounds in terms of the Kullback-Leibler divergence, $K(P, Q)$,

$$mK(P, Q) \leq C(P, Q) \leq MK(P, Q) \quad (3.9)$$

and

$$mK(Q, P) \leq C_\rho(P, Q) - C(P, Q) \leq MK(Q, P). \quad (3.10)$$

(iii) The case $s = 1/2$ yields the information bounds in terms of the Hellinger's discrimination, $h(P, Q)$,

$$4mh(P, Q) \leq C(P, Q) \leq 4Mh(P, Q) \quad (3.11)$$

and

$$4m\left(\frac{1}{4}\eta_{1/2}(P, Q) - h(P, Q)\right) \leq C_\rho(P, Q) - C(P, Q) \leq 4M\left(\frac{1}{4}\eta_{1/2}(P, Q) - h(P, Q)\right). \quad (3.12)$$

(iv) For $s = 0$, the information bounds in terms of the Kullback-Leibler and χ^2 -divergences,

$$mK(P, Q) \leq C(P, Q) \leq MK(P, Q) \quad (3.13)$$

and

$$m(\chi^2(Q, P) - K(Q, P)) \leq C_\rho(P, Q) - C(P, Q) \leq M(\chi^2(Q, P) - K(Q, P)). \quad (3.14)$$

4. NEW INFORMATION DIVERGENCE MEASURE

Let the convex function $f : (0, \infty) \rightarrow \mathbb{R}$ be

$$f(u) = \frac{(u+1)(u-1)^2}{u} \ln \frac{u+1}{2\sqrt{u}}. \quad (4.1)$$

Then, we have the following new divergence measure belonging to the Csiszár's f -divergence family,

$$S(P, Q) = \sum \frac{(p+q)(p-q)^2}{pq} \ln \frac{p+q}{2\sqrt{pq}}. \quad (4.2)$$

Since we can express $S(P, Q)$ as

$$S(P, Q) = \sum \left[\frac{(p+q)(p-q)^2}{pq} \right] \ln \left[\frac{p+q}{2} \cdot \frac{1}{\sqrt{pq}} \right],$$

this measure can be termed as the *Symmetric Chi-Square*, *Arithmetic*, and *Geometric Mean* divergence measure.

It may be noted that, $f(u)$ in (4.1) satisfies $f(1) = 0$, so that $S(P, P) = 0$. Convexity of $f(u)$ ensures that divergence measure $S(P, Q)$ is nonnegative. Thus, we have

- (a) $S(P, Q) \geq 0$ and $S(P, Q) = 0$, iff $P = Q$.
- (b) $S(P, Q)$ is symmetric with respect to probability distribution.
- (c) Since $f^*(u) \equiv uf(1/u) = f(u)$, function $f(u)$ is the $*$ -self conjugate. Therefore, Properties (i)-(vi) of Section 2 hold good for $f(u)$.

We now derive information divergence inequalities providing bounds for $S(P, Q)$ in terms of the well known divergence measures in the following propositions.

PROPOSITION 4.1. Let $S(P, Q)$ and $\Delta(P, Q)$ be defined as (4.2) and (2.7), respectively. Then, inequality

$$SM(P, Q) \leq 4 \sum \frac{(p-q)^2}{\sqrt{pq}} - 2\Delta(P, Q). \tag{4.3}$$

PROOF. Consider the Arithmetic (AM), Geometric(GM) and Harmonic mean(HM) inequality, i.e., $HM \leq GM \leq AM$. Then,

$$\begin{aligned} & HM \leq AM \\ \text{or, } & \frac{2pq}{p+q} \leq \frac{p+q}{2} \\ \text{or, } & \ln \frac{p+q}{2\sqrt{pq}} \geq \ln \frac{2\sqrt{pq}}{p+q}. \end{aligned} \tag{4.4}$$

Multiplying both sides of (4.4) by $(p+q)(p-q)^2/pq$, we have

$$\frac{(p+q)(p-q)^2}{pq} \ln \frac{p+q}{2\sqrt{pq}} \geq \frac{(p+q)(p-q)^2}{pq} \ln \frac{2\sqrt{pq}}{p+q}. \tag{4.5}$$

From $HM \leq GM$, we have $2\sqrt{pq}/p+q \leq 1$, and thus,

$$\ln \frac{2\sqrt{pq}}{p+q} = \ln \left(1 + \left(\frac{2\sqrt{pq}}{p+q} - 1 \right) \right) \approx \frac{4\sqrt{pq}}{p+q} - \frac{2pq}{(p+q)^2} - \frac{3}{2}. \tag{4.6}$$

Now, from (4.5), (4.6), and summing over all $x \in \Omega$, we get

$$\begin{aligned} \sum \frac{(p+q)(p-q)^2}{pq} \ln \frac{p+q}{2\sqrt{pq}} &\leq 4 \sum \frac{(p-q)^2}{\sqrt{pq}} - 2 \sum \frac{(p-q)^2}{p+q}, \\ SM(P, Q) &\leq 4 \sum \frac{(p-q)^2}{\sqrt{pq}} - 2\Delta(P, Q), \end{aligned}$$

and hence, the proof.

Next, we derive the information bounds in terms of the χ^2 -divergence, that is, $\chi^2(P, Q)$.

PROPOSITION 4.2. Let $\chi^2(P, Q)$ and $S(P, Q)$ be defined as (2.3) and (4.2), respectively. For $P, Q \in \mathbb{P}^2$ and $0 < r \leq p/q \leq R < \infty$, we have,

$$\begin{aligned} 0 &\leq S(P, Q) \\ &\leq \left(\frac{1}{4r^3(r+1)} \right) \left[4(r^2-r+1)(r+1)^2 \ln \frac{r+1}{2\sqrt{r}} + (3r^2+4r+3)(r-1)^2 \right] \chi^2(P, Q) \end{aligned} \tag{4.7}$$

and

$$\begin{aligned} 0 &\leq S_\rho(P, Q) - S(P, Q) \\ &\leq \left(\frac{1}{4r^3(r+1)} \right) \left[4(r^2-r+1)(r+1)^2 \ln \frac{r+1}{2\sqrt{r}} + (3r^2+4r+3)(r-1)^2 \right] \chi^2(P, Q), \end{aligned} \tag{4.8}$$

where

$$S_\rho(P, Q) = \sum \frac{(p-q)^2}{2p^2q} \left[2(2p^2+pq+q^2) \ln \frac{p+q}{2\sqrt{pq}} + p^2 - 2pq + q \right]. \tag{4.9}$$

PROOF. From the expression of $f(u)$ in (4.1), we have

$$f'(u) = \left(\frac{u-1}{2u^2} \right) \left[2(2u^2+u+1) \ln \frac{u+1}{2\sqrt{u}} + (u-1)^2 \right], \tag{4.10}$$

and thus,

$$S_\rho(P, Q) = \sum (p - q) f' \left(\frac{p}{q} \right) = \sum \left(\frac{(p - q)^2}{2p^2q} \right) \left[2(2p^2 + pq + q^2) \ln \frac{p + q}{2\sqrt{pq}} + p^2 - 2pq + q \right].$$

Further,

$$f''(u) = \left(\frac{1}{2u^3(u+1)} \right) \left[4(u^2 - u + 1)(u + 1)^2 \ln \frac{u+1}{2\sqrt{u}} + (3u^2 + 4u + 3)(u - 1)^2 \right]. \quad (4.11)$$

Now, if $u \in [r, R] \subset (0, \infty)$, then

$$0 \leq f''(u) \leq \left(\frac{1}{2r^3(r+1)} \right) \left[4(r^2 - r + 1)(r + 1)^2 \ln \frac{r+1}{2\sqrt{r}} + (3r^2 + 4r + 3)(r - 1)^2 \right], \quad (4.12)$$

where r and R are defined above. In view of (3.7), (3.8), and (4.12), we get inequalities (4.7) and (4.8), respectively.

Now, the information bounds in terms of the Kullback-Leibler divergence, $K(P, Q)$ follows.

PROPOSITION 4.3. *Let $K(P, Q)$, $S(P, Q)$, and $S_\rho(P, Q)$ be defined as (2.6), (4.2), and (4.9), respectively. If $P, Q \in \mathbb{P}^2$ and $0 < r \leq p/q \leq R < \infty$, then,*

$$0 \leq S(P, Q) \leq \left(\frac{1}{2r^2(r+1)} \right) \left[4(r^2 - r + 1)(r + 1)^2 \ln \frac{r+1}{2\sqrt{r}} + (3r^2 + 4r + 3)(r - 1)^2 \right] K(P, Q), \quad (4.13)$$

and

$$0 \leq S_\rho(P, Q) - S(P, Q) \leq \left(\frac{1}{2r^2(r+1)} \right) \left[4(r^2 - r + 1)(r + 1)^2 \ln \frac{r+1}{2\sqrt{r}} + (3r^2 + 4r + 3)(r - 1)^2 \right] K(Q, P). \quad (4.14)$$

PROOF. Consider $f''(u)$ as given in (4.11) and let the function $g : [r, R] \rightarrow \mathbb{R}$ be such that,

$$g(u) = u f''(u) = \left(\frac{1}{2u^2(u+1)} \right) \left[4(u^2 - u + 1)(u + 1)^2 \ln \frac{u+1}{2\sqrt{u}} + (3u^2 + 4u + 3)(u - 1)^2 \right]. \quad (4.15)$$

Then,

$$0 \leq g(u) \leq \left(\frac{1}{2r^2(r+1)} \right) \left[4(r^2 - r + 1)(r + 1)^2 \ln \frac{r+1}{2\sqrt{r}} + (3r^2 + 4r + 3)(r - 1)^2 \right]. \quad (4.16)$$

The inequalities (4.13) and (4.14) follow from (3.9), (3.10), and (4.16).

The following proposition provides the information bounds in terms of the Hellinger's discrimination, $h(P, Q)$ and $\eta_{1/2}(P, Q)$.

PROPOSITION 4.4. *Let $h(P, Q)$, $\eta_{1/2}(P, Q)$, $S(P, Q)$, and $S_\rho(P, Q)$ be defined as (2.26), (3.6), (4.2), and (4.11), respectively. For $P, Q \in \mathbb{P}^2$ and $0 < r \leq p/q \leq R < \infty$,*

$$0 \leq S(P, Q) \leq \left(\frac{2}{R^{3/2}(u+1)} \right) \left[4(R^2 - R + 1)(R + 1)^2 \ln \frac{R+1}{2\sqrt{R}} + (3R^2 + 4R + 3)(R - 1)^2 \right] h(P, Q) \quad (4.17)$$

and

$$\begin{aligned}
 0 &\leq S_\rho(P, Q) - S(P, Q) \\
 &\leq \left(\frac{2}{R^{3/2}(u+1)} \right) \left[4(R^2 - R + 1)(R + 1)^2 \ln \frac{R+1}{2\sqrt{R}} + (3R^2 + 4R + 3)(R - 1)^2 \right] \\
 &\quad \times \left(\frac{1}{4} \eta_{1/2}(P, Q) - h(P, Q) \right). \tag{4.18}
 \end{aligned}$$

PROOF. For $f(u)$ in (4.1), we have $f''(u)$ given by (4.11). Let the function $g : [r, R] \rightarrow \mathbb{R}$ be such that

$$\begin{aligned}
 g(u) &= u^{3/2} f''(u) \\
 &= \left(\frac{1}{2u^{3/2}(u+1)} \right) \left[4(u^2 - u + 1)(u + 1)^2 \ln \frac{u+1}{2\sqrt{u}} + (3u^2 + 4u + 3)(u - 1)^2 \right], \tag{4.19}
 \end{aligned}$$

then,

$$\begin{aligned}
 0 &\leq g(u) \\
 &\leq \left(\frac{1}{2R^{3/2}(u+1)} \right) \left[4(R^2 - R + 1)(R + 1)^2 \ln \frac{R+1}{2\sqrt{R}} + (3R^2 + 4R + 3)(R - 1)^2 \right]. \tag{4.20}
 \end{aligned}$$

Thus, inequalities (4.17) and (4.18) are established using (3.11), (3.12), and (4.20).

Next, follows the information bounds in terms of the Kullback-Leibler and χ^2 -divergences.

PROPOSITION 4.5. Let $\chi^2(P, Q)$, $K(P, Q)$, $S(P, Q)$, and $S_\rho(P, Q)$ be defined as (2.3), (2.6), (4.2), and (4.11), respectively. If $P, Q \in \mathbb{P}^2$ and $0 < r \leq p/q \leq R < \infty$, then,

$$\begin{aligned}
 0 &\leq S(P, Q) \\
 &\leq \left(\frac{1}{2R(R+1)} \right) \left[4(R^2 - R + 1)(R + 1)^2 \ln \frac{R+1}{2\sqrt{R}} + (3R^2 + 4R + 3)(R - 1)^2 \right] K(P, Q) \tag{4.21}
 \end{aligned}$$

and

$$\begin{aligned}
 0 &\leq S_\rho(P, Q) - S(P, Q) \\
 &\leq \left(\frac{1}{2R(R+1)} \right) \left[4(R^2 - R + 1)(R + 1)^2 \ln \frac{R+1}{2\sqrt{R}} + (3R^2 + 4R + 3)(R - 1)^2 \right] \\
 &\quad \times (\chi^2(Q, P) - K(Q, P)). \tag{4.22}
 \end{aligned}$$

PROOF. From the expression (4.1), we have $f''(u)$ as given in (4.11). Let the function $g : [r, R] \rightarrow \mathbb{R}$ be such that,

$$\begin{aligned}
 g(u) &= u^2 f''(u) \\
 &= \left(\frac{1}{2u(u+1)} \right) \left[4(u^2 - u + 1)(u + 1)^2 \ln \frac{u+1}{2\sqrt{u}} + (3u^2 + 4u + 3)(u - 1)^2 \right], \tag{4.23}
 \end{aligned}$$

then,

$$\begin{aligned}
 0 &\leq g(u) \\
 &\leq \left(\frac{1}{2R(R+1)} \right) \left[4(R^2 - R + 1)(R + 1)^2 \ln \frac{R+1}{2\sqrt{R}} + (3R^2 + 4R + 3)(R - 1)^2 \right]. \tag{4.24}
 \end{aligned}$$

Thus, (4.21) and (4.22) follow from (3.13), (3.14), and (4.24).

5. PARAMETRIC MEASURE OF INFORMATION

The parametric measures of information are applicable to regular families of probability distributions, that is, to the families for which the following regularity conditions are assumed to be satisfied. Let for $\theta = (\theta_1, \dots, \theta_k)$, the Fisher information matrix [2] be

$$I_x(\theta) = \begin{cases} E_\theta \left[\frac{\partial}{\partial \theta} \log f(X, \theta) \right]^2, & \text{if } \theta \text{ is univariate,} \\ \left\| E_\theta \left[\frac{\partial}{\partial \theta_i} \log f(X, \theta) \frac{\partial}{\partial \theta_j} \log f(X, \theta) \right] \right\|_{k \times k}, & \text{if } \theta \text{ is } k\text{-variate,} \end{cases} \quad (5.1)$$

where $\|_{k \times k}$ denotes a $k \times k$ matrix.

The regularity conditions are,

- (R1) $f(x, \theta) > 0$, for all $x \in \Omega$ and $\theta \in \Theta$,
- (R2) $\frac{\partial}{\partial \theta_i} f(X, \theta)$ exists, for all $x \in \Omega$ and $\theta \in \Theta$ and all $i = 1, \dots, k$,
- (R3) $\frac{d}{d\theta_i} \int_A f(x, \theta) d\mu = \int_A \frac{d}{d\theta_i} f(x, \theta) d\mu$, for any $A \in \mathcal{A}$ (measurable space (X, \mathcal{A}) in respect of a finite or σ -finite measure μ), all $\theta \in \Theta$, and all i .

In [1], authors suggested the following method to construct the parametric measure from the nonparametric measure.

Let $k(\theta)$ be a one-to-one transformation of the parameter space Θ onto itself, with $k(\theta) \neq \theta$. The quantity

$$I_x[\theta, k(\theta)] = I_x[f(x, \theta), f(x, k(\theta))], \quad (5.2)$$

can be considered as a parametric measure of information based on $k(\theta)$.

This method is employed to construct the modified Csiszár's measure of information about univariate θ contained in X and based on $k(\theta)$ as,

$$I_x^C[\theta, k(\theta)] = \int f(x, \theta) \phi \left(\frac{f(x, k(\theta))}{f(x, \theta)} \right) d\mu. \quad (5.3)$$

Now, we have the following proposition for the parametric measure of information from $S(P, Q)$.

PROPOSITION 5.1. *Let the convex function $\phi : (0, \infty) \rightarrow \mathbb{R}$ be*

$$\phi(u) = \frac{(u+1)(u-1)^2}{u} \ln \frac{u+1}{2\sqrt{u}}, \quad (5.4)$$

and the corresponding nonparametric divergence measure,

$$S(P, Q) = \sum \frac{(p+q)(p-q)^2}{pq} \ln \frac{p+q}{2\sqrt{pq}}.$$

Then, the parametric measure $S^C(P, Q)$ is the same as $S(P, Q)$.

PROOF. For discrete random variables X , the expression (5.3) can be written as

$$I_x^C[\theta, k(\theta)] = \sum_{x \in \Omega} p(x) \phi \left(\frac{q(x)}{p(x)} \right). \quad (5.5)$$

From (5.4), we have

$$\phi \left(\frac{q(x)}{p(x)} \right) = \frac{(p+q)(p-q)^2}{p^2q} \ln \frac{p+q}{2\sqrt{pq}}, \quad (5.6)$$

where we denote $p(x)$, and $q(x)$ by p and q , respectively.

Then, $S^C(P, Q)$ follows from (5.5) and (5.6) as

$$S^C(P, Q) := I_x^C[\theta, k(\theta)] = \sum_{x \in \Omega} \frac{(p+q)(p-q)^2}{pq} \ln \frac{p+q}{2\sqrt{pq}} = S(P, Q), \quad (5.7)$$

and hence, the proposition.

6. NUMERICAL ILLUSTRATION

We consider two examples of symmetrical and asymmetrical probability distributions. We calculate measures $S(P, Q)$, $\Psi(P, Q)$, $\chi^2(P, Q)$, $J(P, Q)$, and verify bounds derived above for $S(P, Q)$.

EXAMPLE 1 SYMMETRICAL. Let P be the binomial probability distribution for the random variable X with parameters $(n = 8, p = 0.5)$ and Q its approximated normal probability distribution. Then, we have Table 1.

Table 1. Binomial Probability Distribution $(n = 8, p = 0.5)$.

x	0	1	2	3	4	5	6	7	8
$p(x)$	0.004	0.031	0.109	0.219	0.274	0.219	0.109	0.031	0.004
$q(x)$	0.005	0.030	0.104	0.220	0.282	0.220	0.104	0.030	0.005
$p(x)/q(x)$	0.774	1.042	1.0503	0.997	0.968	0.997	1.0503	1.042	0.774

The measures $S(P, Q)$, $\Psi(P, Q)$, $\chi^2(P, Q)$, and $J(P, Q)$ are,

$$\begin{aligned}
 S(P, Q) &= 0.00001030, & \Psi(P, Q) &= 0.00305063, \\
 \chi^2(P, Q) &= 0.00145837, & J(P, Q) &= 0.00151848.
 \end{aligned}$$

It is noted that,

$$r (= 0.774179933) \leq \frac{p}{q} \leq R (= 1.050330018).$$

The upper bound for $S(P, Q)$ based on $\chi^2(P, Q)$ divergence from (4.7),

$$\begin{aligned}
 \text{Upper Bound} &= \left(\frac{1}{4r^3(r+1)} \right) \left[4(r^2-r+1)(r+1)^2 \ln \frac{r+1}{2\sqrt{r}} + (3r^2+4r+3)(r-1)^2 \right] \chi^2(P, Q) \\
 &= 0.00158357,
 \end{aligned}$$

and, thus, $0 < S(P, Q) = 0.00001134 < 0.001575814$. The length of the interval is 0.001575814.

The upper bound for $S(P, Q)$ based on $K(P, Q)$ from (4.13),

$$\begin{aligned}
 \text{Upper Bound} &= \left(\frac{1}{2r^2(r+1)} \right) \left[4(r^2-r+1)(r+1)^2 \ln \frac{r+1}{2\sqrt{r}} + (3r^2+4r+3)(r-1)^2 \right] K(P, Q) \\
 &= 0.002850183,
 \end{aligned}$$

and therefore, $0 < S(P, Q) = 0.00001134 < 0.002850183$. The length of the interval is 0.002850183.

EXAMPLE 2 ASYMMETRICAL. Let P be the binomial probability distribution for the random variable X , with parameters $(n = 8, p = 0.4)$ and Q its approximated normal probability distribution. Then, we have Table 2.

Table 2. Binomial Probability Distribution $(n = 8, p = 0.4)$.

x	0	1	2	3	4	5	6	7	8
$p(x)$	0.017	0.090	0.209	0.279	0.232	0.124	0.041	0.008	0.001
$q(x)$	0.020	0.082	0.198	0.285	0.244	0.124	0.037	0.007	0.0007
$p(x)/q(x)$	0.850	1.102	1.056	0.979	0.952	1.001	1.097	1.194	1.401

The measures $S(P, Q)$, $\Psi(P, Q)$, $\chi^2(P, Q)$, and $J(P, Q)$ are,

$$\begin{aligned}
 \Psi M(P, Q) &= 0.00001134, & \Psi(P, Q) &= 0.00305063, \\
 \chi^2(P, Q) &= 0.00145837, & J(P, Q) &= 0.00151848.
 \end{aligned}$$

It is noted that,

$$r (= 0.774179933) \leq \frac{p}{q} \leq R (= 1.050330018).$$

The upper bound for $S(P, Q)$ based on $\chi^2(P, Q)$ divergence from (4.7),

$$\begin{aligned} \text{Upper Bound} &= \left(\frac{1}{4r^3(r+1)} \right) \left[4(r^2-r+1)(r+1)^2 \ln \frac{r+1}{2\sqrt{r}} + (3r^2+4r+3)(r-1)^2 \right] \chi^2(P, Q) \\ &= 0.001575814, \end{aligned}$$

and thus, $0 < S(P, Q) = 0.00001134 < 0.001575814$. The length of the interval is 0.001575814.

The upper bound for $S(P, Q)$ based on $K(P, Q)$ from (4.13),

$$\begin{aligned} \text{Upper Bound} &= \left(\frac{1}{2r^2(r+1)} \right) \left[4(r^2-r+1)(r+1)^2 \ln \frac{r+1}{2\sqrt{r}} + (3r^2+4r+3)(r-1)^2 \right] K(P, Q) \\ &= 0.007996539, \end{aligned}$$

and $0 < S(P, Q) = 0.00001134 < 0.007996539$. The length of the interval is 0.007996539.

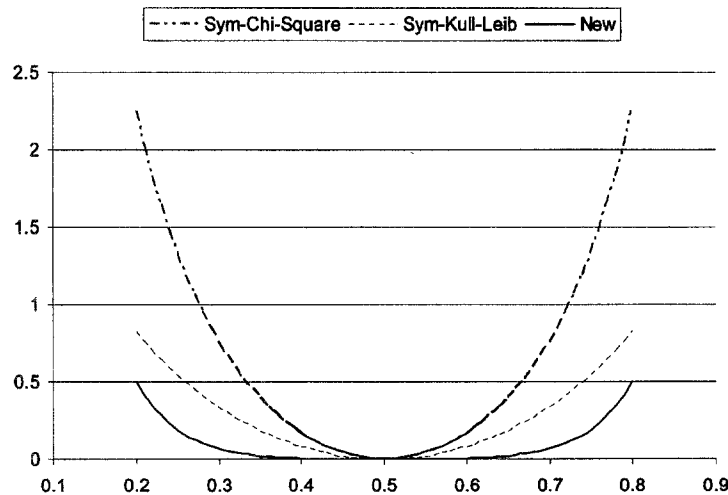


Figure Figure 1. Measures $S(P, Q)$ -New, $\Psi(P, Q)$ -Sym Chi Square, and $J(P, Q)$ -Sym Kullback Leibler.

Figure 1 shows the behavior of $S(P, Q)$ -[New], $\Psi(P, Q)$ -[Sym-Chi-square] and $J(P, Q)$ -[Sym-Kull-Leib]. We have considered $p = (a, 1 - a)$ and $q = (1 - a, a)$, $a \in [0, 1]$. It is clear from Figure 1 that measures $\Psi(P, Q)$ and $J(P, Q)$ have a steeper slope than $S(P, Q)$.

7. CONCLUDING REMARKS

The Csiszár's f -divergence is a general class of divergence measures which includes several divergences used in measuring the distance or affinity between two probability distributions. This class is introduced by using a convex function f , defined on $(0, \infty)$. An important property of this divergence is that many known divergences can be obtained from this measure by appropriately defining the convex function f . Nonparametric measures for the Csiszár's f -divergences are also available. For this class of divergences, its properties, bounds, and relations among well known divergences have been of interest to the researchers. We have introduced a new symmetric divergence measure by considering a convex function and have investigated its properties. Further, we have established its bounds in terms of known divergence measures. Work on one parametric generalization of this measure is in progress and will be reported elsewhere.

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